Algorithms for context prediction in Ubiquitous Systems

Markov prediction approaches

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Overview and Structure

- Introduction to context aware computing
- Basics of probability theory
- Algorithms
 - Simple prediction approaches: ONISI and IPAM
 - Markov prediction approaches
 - The State predictor
 - Alignment prediction
 - Prediction with self organising maps
 - Stochastic prediction approaches: ARMA and Kalman filter
 - Alternative prediction approahces
 - Dempster shafer
 - Evolutionary algorithms
 - Neural networks
 - Simulated annealing

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Outline

Markov prediction approaches

- Introduction and Markov properties
- 2 Markov chains
- 8 Hidden Markov Models
- 4 Context prediction with Markov approaches
 - Properties of Markov prediction approaches
- Conditional random fields
 - Context prediction with CRF
 - Properties of CRF prediction approaches

6 Conclusion

Historical remarks

Markov processes

- Intensively studied
- Major branch in the theory of stochastic processes
- A. A. Markov (1856 1922)
- Extended by A. Kolmogorov by chains of infinitely many states
 - 'Anfangsgründe der Theorie der Markoffschen Ketten mit unendlich vielen möglichen Zuständen' (1936)¹

¹A. Kolmogorov, Anfangsgründe der Theorie der Markoffschen Ketten mit unendlich vielen möglichen Zuständen, 1936.

Historical remarks

Markov Chains

- Theory of Markov chains applied to a variety of algorithmic problems
- Standard tool in many probabilistic applications
- Intuitive graphical representation
 - Possible to illustrate properties of stochastic processes graphically
- Popular for their simplicity and easy applicability to huge set of problems²

²William Feller, *An introduction to probability theory and its applications*, Wiley, 1968. Stephan Sigg Algorithms for context prediction in Ubiquitous Systems 6/111

Introcution

- Dependent trials of events
 - Set of possible outcomes of a measurement *E_i* associated with occurence probability *p_i*
- When occurence of events is not independent
 - Probability to observe specific sequence E_1, E_2, \ldots, E_i obtained by conditional probability:

$$P(E_i|E_1, E_2, \dots, E_{i-1})$$
 (1)

• In general:

$$P(E_i|E_1, E_2, \dots, E_{i-1}) \neq P(E_i|E_2, \dots, E_{i-1})$$
(2)

Independent random variables

- Sequence of tials for independent random variable
- T: number of trials up to first success of probability p.
- Then:

$$P\{T > k\} = (1 - p)^k$$
(3)

- Suppose: No success during the first *m* trials
- Waiting time T to first success for m-th trial has same distribution $(1-p)^k$
- Independent of number of preceding failures m

Examples

• Independent random variables

- Number of coin tosses until 'head' is observed
- Radioactive atoms always have the same probability of decaying at the next trial
- Dependent random variables
 - The knowledge that no streetcar has passed for five minutes increases our expectation that it will come soon.
 - Coin tossing:
 - Probability that the cumulative numbers of heads and tails will equalize at the second trial is $\frac{1}{2}$
 - Given that they did not, the probability that they equalize after two additional trials is only $\frac{1}{4}$

Lack of memory - Rigorous formulation

- Suppose a waiting time *T* assumes the values 0, 1, 2, ... with probabilities *p*₀, *p*₁, *p*₂, ...
- Let T have the following property
 - Conditional probability that the waiting time terminates at the k-th trial equals p_0
- Then:
 - $p_k = (1 p_0)^k p_0$

Lack of Memory - Rigorous formulation

Proof.

- $1 p_k = p_{k+1} + p_{k+2} + \dots = P\{T > k\}$
- Conditional probability of T = k: $p_k/(1 p_{k-1})$
- Assumption for all $k \ge 1$: $\frac{p_k}{1-p_{k-1}} = p_0$
- Since $p_k = (1 p_{k-1}) (1 p_k)$

$$\frac{1-p_k}{1-p_{k-1}} = 1-p_0 \tag{4}$$

• since $1 - p_0 = p_1 + p_2 + \ldots : 1 - p_k = (1 - p_0)^{k+1}$ and

$$p_k = 1 - p_{k-1} - (1 - p_k) = (1 - p_0)^k p_0$$
 (5)

Markov property

Markov property

In the theory of stochasitc processes the described lack of memory is connected with the Markovian property.

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Dependence and independence of events

- Independent trials of events
 - Set of possible outcomes of a measurement *E_i* associated with occurence probability *p_i*
- Probability to observe sample sequence:

•
$$P\{(E_1, E_2, ..., E_i)\} = p_1 p_2 \cdots p_i$$

Dependence and independence of events

- Theory of Markov chains:
 - Outcome of any trial depends exclusively on the outcome of the directly preceding trial
 - Outcome of E_k is no longer associated with fixed probability p_k
 - Instead: With every pair (E_i, E_j) a conditional probability p_{ij}
 - Probability that E_j is observed after E_i
 - Additionally: Probability *a_i* of the event *E_i*

Dependence and independence of events

• Theory of Markov chains:

•
$$P\{(E_i, E_j)\} = a_i p_{ij}$$

• $P\{(E_i, E_j, E_k)\} = a_i p_{ij} p_{jk}$
• $P\{(E_i, E_j, E_k, E_l)\} = a_i p_{ij} p_{jk} p_{kl}$
• $P\{(E_i, E_j, \dots, E_m, E_n)\} = a_i p_{ij} \dots p_{mn}$

Markov chain

Markov chain

A sequence of observations E_1, E_2, \ldots is called a Markov chain if the probabilities of sample sequences are defined by

$$P(E_1, E_2, \dots, E_i) = a_1 \cdot p_{12} \cdot p_{23} \cdot \dots \cdot p_{(i-1)i}.$$
 (6)

and fixed conditional probabilities p_{ij} that the event E_i is observed directly in advance of E_j .

Markov chain

• Markov chain described by probability *a* for initial distribution and matrix *P* of transition probabilities.

$$P = \begin{bmatrix} p_{11} & p_{12} & p_{13} & \cdots \\ p_{21} & p_{22} & p_{23} & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$
(7)

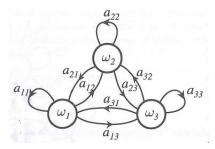
• *P* is a square matrix with non-negative entries that sum to 1 in each row.

Stochastic matrix

- P is called a stochastic matrix.
- Any stochastic matrix is suited to describe transition probabilities of Markov chains.

Markov chain

- Markov chain sometimes modelled as directed graph G = (V, E)
- Labelled edges in E
- states (or vertices) in V.
- Transition probabilities p_{ij} between $E_i, E_j \in V$



Derive state transision probabilities

- *p*^k_{ij} denotes probability that *E_j* is observed exactly *k* observations after *E_i* was observed.
- Calculated as the sum of the probabilities for all possible paths $E_i E_{i_1} \cdots E_{i_{k-1}} E_j$ of length k
- We already know

$$p_{ij}^1 = p_{ij} \tag{8}$$

Consequently:

$$P_{ij}^2 = \sum_{\nu} p_{i\nu} \cdot p_{\nu j} \tag{9}$$

Derive state transision probabilities

• By mathematical induction:

$$p_{ij}^{n+1} = \sum_{\nu} p_{i\nu} \cdot p_{\nu j}^n \tag{10}$$

and

$$p_{ij}^{n+m} = \sum_{\nu} p_{i\nu}^{m} \cdot p_{\nu j}^{n} = \sum_{\nu} p_{i\nu}^{n} \cdot p_{\nu j}^{m}$$
(11)

Derive state transision probabilities

- Similar to the matrix P we can create a matrix Pⁿ that contains all pⁿ_{ii}
- We obtain P_{ij}^{n+1} from P^{n+1} by multiplying all elements of the *i*-th row of *P* with the corresponding elements of the *j*-the column of P^n and add all products.
- Symbolically: $P^{n+m} = P^n P^m$.

Examples

Markov chains:

- Urn models
 - Every Markov chain is equivalent to an urn model
 - Each urn represents a state in a markov chain and probabilities to draw specific balls represent possible events in this state
- Branching processes
 - Instead of saying that the *n*-th trial results in *E_k* we say that the *n*-th generation is of size *k*
- Random walk on a line
 - Events are transitions between states
 - Only two events are possible in each state

Random walks and ruin problems

Random walk

• When there are only two possible states E_1 and E_2 the matrix of transition probabilities is of the form

$$P = \begin{bmatrix} 1-p & p\\ \alpha & 1-\alpha \end{bmatrix}$$
(12)

- Can be realised by particle moving along one axis in one or the other direction.
- System is in state E_1 when the particle moves into one direction and in state E_2 otherwise.

Random walks and ruin problems

• Possible problems / questions

- Expected time to return to origin
- Expected time to return to origin given that the starting point had a specific distance to the origin

• • • •

Random walks and ruin problems

• Random walk with absorbing barriers

$$P = \begin{bmatrix} 1 & 0 & 0 & 0 & \cdots & 0 & 0 & 0 \\ 1-p & 0 & p & 0 & \cdots & 0 & 0 & 0 \\ 0 & 1-p & 0 & p & \cdots & 0 & 0 & 0 \\ \vdots & & & \vdots & & & \\ 0 & 0 & 0 & 0 & \cdots & 1-p & 0 & p \\ 0 & 0 & 0 & 0 & \cdots & 0 & 0 & 1 \end{bmatrix}$$
(13)

- First and last state are obsorbing
- All inner states implement a random walk on the line
- Possible application: Game between two players with equal money balance where the loosing one has to pay one unit to the winner.

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Random walks and ruin problems

Random walk with reflecting barriers

$$P = \begin{bmatrix} 1-p & p & 0 & 0 & \cdots & 0 & 0 & 0 \\ 1-p & 0 & p & 0 & \cdots & 0 & 0 & 0 \\ 0 & 1-p & 0 & p & \cdots & 0 & 0 & 0 \\ \vdots & & \vdots & & & \\ 0 & 0 & 0 & 0 & \cdots & 1-p & 0 & p \\ 0 & 0 & 0 & 0 & \cdots & 0 & 1-p & p \end{bmatrix}$$
(14)

• First and last state are reflecting

• All inner states implement a random walk on the line

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Random walks and ruin problems

Classical ruin problem

- Consider a gambler who wins or loses a dollar with probabilities p and 1 - p
- Initial capital of gambler and adversary: z, a z
- Game ends when the capital reaches 0 or a.
 - When one of the players is ruined
- We are interested in the probability of the gamblers ruin and the probability distribution of the game

Random walks and ruin problems

- Gamblers ruin problem
 - Random walk with absorbing barriers at 0 and a
- Examples:
 - Physicists use this model as crude approximation to one-dimensional diffusion or Brownian motion (Particle is exposed to great number of molecular collisions which impart to it a random motion)
 - p > 1/2 represents a drift to the right, when shocks from the left are more probable

Random walks and ruin problems

- Probability of gamblers ruin
 - *q_z*: Probability of gambler's ultimate ruin when *z* is the starting capital and *a* is the overal capital
 - After the first trial the gablers's fortune is either z 1 or z + 1:

•
$$q_z = pq_{z+1} + (1-p)q_{z-1}$$

• We can show:

$$p \neq \frac{1}{2} \quad : \quad q_z = \frac{\left(\frac{1-p}{p}\right)^a - \left(\frac{1-p}{p}\right)^z}{\left(\frac{1-p}{p}\right)^a - 1} \tag{15}$$
$$p = \frac{1}{2} \quad : \quad q_z = 1 - \frac{z}{a} \tag{16}$$

Random walks and ruin problems

- Probability of gamblers ruin
 - The probability p_z of the gambler winning the game is equal to the probability of his adversary loosing the game.
 - It is therefore obtained in the same way by replacing p with

$$1-p$$
 and z by $a-z$

• Therefore:
$$p_z + q_z = 1$$

Random walks and ruin problems

- Some interesting results
 - Since for $p = \frac{1}{2}$, we have derived $q_z = 1 \frac{z}{a}$
 - A player with initial capital z = 999 has a probability of 0.999 to win a dollar before losing his capital.
 - With p = 0.4 the game is unfavorable, but still the probability of winning a dollar before losing the capital is about $\frac{2}{3}$

Random walks and ruin problems

Example – anecdote

A certain man used to visit Monte Carlo year after year and was always successful in recovering the cost of his vacations. He firmly believed in a magic power over chance.

This experience is not surprising.

- Assuming that he started with ten times the ultimate gain, the chances of success in any year are nearly 0.9.
- The probability of an unbroken sequence of ten successes is about $(1 \frac{1}{10})^{10} \approx e^{-1} \approx 0.37$

Therefore, continued success is by no means improbable

• However: one failue would result in the gambler's ruin :-)

Random walks and ruin problems

р	9	z	а	Probability of		Expected	
				Ruin	Success	Gain	Duration
0.5	0.5	9	10	0.1	0.9	0	9
0.5	0.5	90	100	0.1	0.9	0	900
0.5	0.5	900	1,000	0.1	0.9	0	90,000
0.5	0.5	950	1,000	0.05	0.95	0	47,500
0.5	0.5	8,000	10,000	0.2	0.8	0	16,000,000
0.45	0.55	9	10	0.210	0.790	-1.1	11
0.45	0.55	90	100	0.866	0.134	-76.6	765.6
0.45	0.55	99	100	0.182	0.818	-17.2	171.8
0.4	0.6	90	100	0.983	0.017	-88.3	441.3
0.4	0.6	99	100	0.333	0.667	-32.3	161.7

 Effect of increasing stakes is more pronounced than might be expected

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Random walks and ruin problems

- Expected duration of the game
 - *D_z*: Expected duration of the game when *z* is the starting capital and *a* is the overal capital
 - After the first trial the gablers's fortune is either z 1 or z + 1:

•
$$D_z = pD_{z+1} + (1-p)D_{z-1} + 1$$

We can show:

$$p \neq \frac{1}{2} \quad : \quad D_z = \frac{z}{1 - 2p} - \frac{a}{1 - 2p} \cdot \frac{1 - \left(\frac{1 - p}{p}\right)^z}{1 - \left(\frac{1 - p}{p}\right)^a} \quad (17)$$
$$p = \frac{1}{2} \quad : \quad D_z = z(a - z) \tag{18}$$

Random walks and ruin problems

• Expected duration of the game

$$p \neq \frac{1}{2} \quad : \quad D_z = \frac{z}{1 - 2p} - \frac{a}{1 - 2p} \cdot \frac{1 - \left(\frac{1 - p}{p}\right)^z}{1 - \left(\frac{1 - p}{p}\right)^a} \quad (19)$$
$$p = \frac{1}{2} \quad : \quad D_z = z(a - z) \quad (20)$$

Examples – Duration considerably longer as naively expected:

- If two players with 500 dollars each toss a fair coin, average duration of the game is 250000 trials
- If a gambler has only one dollar and his adversary 1000, the average duration is 1000 trials

Closures and closed sets

Closed set of states

- A set *C* of states is closed if no state outside *C* can be reached from any state *E_i* in *C*.
- For an arbitrary set *C* of states the smallest closest set containing *C* is called the closure of *C*
- A single state E_k forming a closed set is called absorbing

Closures and closed sets

Closed sets in stochastic matrices

If in a matrix P^n all rows and all columns corresponding to states outside a closed set C are deleted, the remaining matrices are again stochastic matrices.

Closures and closed sets

Irreducible Markov chain

A Markov chain is irreducible if there exists no closed set other than the set of all states.

Criterion for irreducible chains

A chain is irreducible if, and only if, every state can be reached from every other state.

Periodicity

Periodicity of states

- The state E_j has period t > 1 if $p_{jj}^n = 0$ unless n = vt is a multiple of t and t is the largest integer with this property.
- the state E_i is aperiodic if no such t > 1 exists

• A state E_i to which no return i possible is considered aperiodic

Periodicity

- To deal with a periodic *E_j* it suffices to consider the chain at the trials *t*, 2*t*, 3*t*
- In this way we obtain a new Markov chain with transition probabilities p^t_{ik}
- In this new chain E_j is aperiodic
- Results concerning aperiodic states can thus be transferred to periodic states

Persistent and transient states

Persistent and transient states

• The state E_j is persistent if $\sum_{n=1}^{\infty} p_{jj}^n = 1$ and transient if $\sum_{n=1}^{\infty} p_{jj}^n < 1$

• A persistent state E_j is called null state if its mean recurrence time $\mu_j = \infty$

Irreducible chains

- Two states are of the same type when they are either
 - both aperiodic
 - both have the same period
 - both are transient
 - both are persistent and each
 - with infinite recurrence times
 - or finite recurrence times

Irreducible chains

Type of states in irreducible chains

All states of an irreducible chain are of the same type

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Algorithms for context prediction in Ubiquitous Systems

Applications – Card shuffling

- A deck of *N* cards can be arranged in *N*! different orders.
- Each order represent a possible state of the system
- We conceive each particular shuffling operation as a transformation $E_i \rightarrow E_j$
- Result:
 - The permutation is not cyclic
 - Therefore, repeated application of a single operation will never visit all possible states
 - This means that the original state is again observed before all states are visited
- This is a Markov chain:
 - We assume that a player applies several shuffling operation with a random probability and that the current order of the cards is not known.

Markov processes

Markov process

A sequence of discrete-valued random variables is a markov process if the joint distribution of (X^1, \ldots, X^n) is defined in such a way that the conditional probability of the relation $X^n = x$ on the hypothesis $X^{n_1} = x_1, \ldots, X^{n_r} = x_r$ is identical with the conditional probability of $X^n = x$ on the single hypothesis $X^{n_r} = x_r$.

Higher order Markov processes

- Order k Markov processes
- Typically
 - Occurence of event dependent on *k* events that were observed directly beforehand
 - Constrained lack of memory
 - Dependence between the last k events observed
- Useful for context prediction / time series forecasting, when typical patterns or trends are to be considered

Higher order Markov processes

• Probability that E_1, E_2, \ldots, E_i observed is then

$$P(E_1, E_2, \dots, E_i) = p_1 \cdot p_{12} \cdot p_{23} \cdot \dots \cdot p_{(i-1)i}.$$
(21)

• Required: $p_i > 0 \forall i \text{ and } \sum_{p_i} = 1.$

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Introduction

- Make a sequence of decisions for a process that is not directly observable³
- Current states of the process might be impacted by prior states
- HMM often utilised in speach recognition or gesture recognition

³Richard O. Duda, Peter E. Hart and David G. Stork, *Pattern classification*, Wiley interscience, 2001. Stephan Sigg Algorithms for context prediction in Ubiquitous Systems

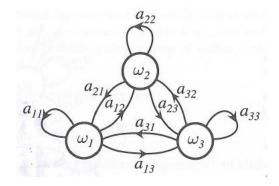
Applications

- Computational biology
 - Align biological sequences
 - Find sequences homologous to a known evolutionary family
 - Analyse RNA secondary structure ⁴
- Computational linguistics⁵
 - Topic segmentation of text
 - Information extraction

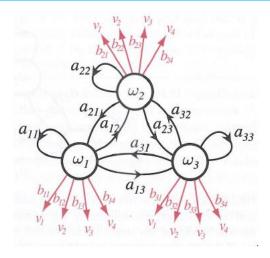
⁴R. Durbin, S. Eddy, A. Krogh and G. Mitchison, *Biological sequence analysis: Probabilistic models of proteins and nucleic acids*, Cambridge University Press, 1998.

⁵C.D. Manning and H. Schütze, *Foundations of statistical natural language processing*, MIT Press, 1999. Stephan Sigg Algorithms for context prediction in Ubiquitous Systems

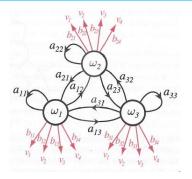
First order Markov models



First order hidden Markov models



First order hidden Markov models



- At every time step t the system is in an internal state $\omega(t)$
- Additionally, we assume that it emits a (visible) symbol v(t)
- Only access to visible symbols and not to intenal states

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First order hidden Markov models

•
$$V^T = \{v(1), v(2), \dots, v(T)\}$$

- In any state ω(t) we have a probability of emitting a particular visible symbol v_k(t)
- Probability to be in state ω_j(t) and emit symbol v_k(t):
 P(v_k(t)|ω_j(t)) = b_{jk}
- Transmission probabilities: $p_{ij} = P(\omega_j(t+1)|\omega_i(t))$
- Emission probability: $b_{jk} = P(v_k(t)|\omega_j(t))$

First order hidden Markov models

Normalisation conditions

$$\sum_{j} p_{ij} = 1 \text{ for all } i$$

$$\sum_{k} b_{jk} = 1 \text{ for all } j$$
(22)
(23)

First order hidden Markov models

• Central issues in hidden Markov models:

- Evaluation problem
 - Determine the probability that a particular sequence of visible symbols V^T was generated by a given hidden Markov model
- Decoding problem
 - Determine the most likely sequence of hidden states ω^T that led to a specific sequence of observations V^T
- Learning problem
 - Given a set of training observations of visible symbols, determine the parameters p_{ij} and b_{jk} for a given HMM

Ρ

First order hidden Markov models - Evaluation problem

• Probability that model produces a sequence V^{T} :

$$P(V^{T}) = \sum_{\overline{\omega}^{T}} P(V^{T} | \overline{\omega}^{T}) P(\overline{\omega}^{T})$$
(24)

Also:

$$P(\overline{\omega}^{T}) = \prod_{t=1}^{T} P(j\omega(t)|\omega(t-1))$$
(25)
$$(V^{T}|\overline{\omega}^{T}) = \prod_{t=1}^{T} P(v(t)|\omega(t))$$
(26)

• Together:

$$P(V^{T}) = \sum_{\overline{\omega}^{T}} \prod_{t=1}^{T} P(v(t)|\omega(t)) P(\omega(t)|\omega(t-1))$$
(27)

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First order hidden Markov models - Evaluation problem

• Probability that model produces a sequence V^T :

$$P(V^{T}) = \sum_{\overline{\omega}^{T}} \prod_{t=1}^{T} P(v(t)|\omega(t)) P(\omega(t)|\omega(t-1))$$
(28)

- Formally complex but straightforward
- Naive computational complexity
 O(c^TT)

First order hidden Markov models - Evaluation problem

• Probability that model produces a sequence V^T :

$$P(V^{T}) = \sum_{\overline{\omega}^{T}} \prod_{t=1}^{T} P(v(t)|\omega(t))P(\omega(t)|\omega(t-1))$$
(29)

- Computationally less complex algorithm:
 - Calculate $P(V^T)$ recursively
 - $P(v(t)|\omega(t))P(\omega(t)|\omega(t-1))$ involves only $v(t), \omega(t)$ and $\omega(t-1)$

$$\alpha_{j}(t) = \begin{cases} 0 & t = 0 \text{ and } j \neq \text{ initial state} \\ 1 & t = 0 \text{ and } j = \text{ initial state} \\ \left[\sum_{i} \alpha_{i}(t-1)p_{ij}\right] b_{jk}v(t) & \text{otherwise} \end{cases}$$
(30)

First order hidden Markov models - Evaluation problem

Forward Algorithm

• Computational complexity: $O(c^2 T)$

Forward algorithm

1 initialise
$$t \leftarrow 0, p_{ij}, b_{jk}, V^T, \alpha_j(o)$$

2 for $t \leftarrow t+1$
3 $\alpha_j(t) \leftarrow b_{jk}v(t) \sum_{i=1}^c \alpha_i(t-1)p_{ij}$
4 until $t = T$
5 return $P(V^T) \leftarrow \alpha_0(T)$ for the final state
6 end

First order hidden Markov models - Decoding problem

- Given a sequence V^T, find the most probable sequence of hidden states.
- Enumeration of every possible path will cost $O(c^T T)$
 - Not feasible

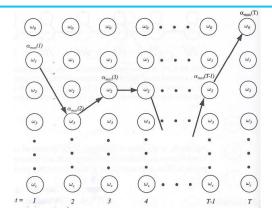
First order hidden Markov models - Decoding problem

• Given a sequence V^T , find the most probable sequence of hidden states.

Decoding algorithm

```
initialise path \leftarrow {}, t \leftarrow 0
1
2
         for t \leftarrow t+1
3
             i \leftarrow i+1
4
             for i \leftarrow i+1
                 \alpha_i(t) \leftarrow b_{ik}v(t)\sum_{i=1}^{c} \alpha_i(t-1)p_{ii}
5
6
             until j = c
7
             j' \leftarrow \arg \max_i \alpha_i(t)
8
             append \omega_{i'} to path
9
         until t = T
10
    return path
11 end
```

First order hidden Markov models - Decoding problem



- Computational time of the decoding algorithm
 O(c²T)
- However, computed path might be invalid

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First order hidden Markov models - Learning problem

- Determine the model parameters p_{ij} and b_{jk}
 - Given: Training sample of observed values V^T
- No method known to obtain the optimal or most likely set of parameters from the data
 - However, we can nearly always determine a good solution by the forward-backward algorithm
 - General expectation maximisation algorithm
 - Iteratively update weights in order to better explain the observed training sequences

First order hidden Markov models - Learning problem

 Probability that the model is in state ω_i(t) and will generate the remainder of the given target sequence:

$$\beta_{i}(t) = \begin{cases} 0 & t = T \text{ and } \omega_{i}(t) \neq \omega_{0} \\ 1 & t = T \text{ and } \omega_{i}(t) = \omega_{0} \\ \sum_{j} \beta_{j}(t+1)p_{ij}b_{jk}v(t+1) & \text{otherwise} \end{cases}$$
(31)

First order hidden Markov models - Learning problem

- α_i(t) and β_i(t) only estimates of their true values since transition probabilities p_{ij}, b_{jk} unknown
- Probability of transition between $\omega_i(t-1)$ and $\omega_j(t)$ can be estimated
 - Provided that the model generated the entire training sequence V^T by any path

$$\gamma_{ij}(t) = \frac{\alpha(t-1)\rho_{ij}b_{jk}\beta_j(t)}{P(V^T|\Theta)}$$
(32)

• Probability that model generated sequence V^T :

$$P(V^{T}|\Theta) \tag{33}$$

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First order hidden Markov models - Learning problem

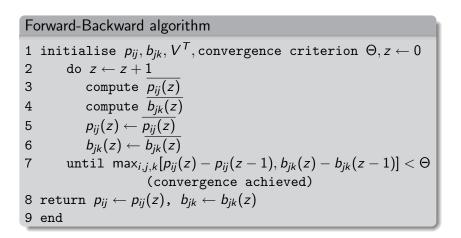
• Calculate improved estimate for p_{ij} and b_{jk}

$$\overline{p_{ij}} = \frac{\sum_{t=1}^{T} \gamma_{ij}(t)}{\sum_{t=1}^{T} \sum_{k} \gamma_{ik}(t)}$$
(34)

$$\overline{b_{jk}} = \frac{\sum_{t=1,v(t)=v_k}^T \sum_l \gamma_{jl}(t)}{\sum_{t=1}^T \sum_l \gamma_{jl}(t)}$$
(35)

- Start with rough estimates of p_{ij} and b_{jk}
- Calculate improved estimates
- Repeat until some conversion is reached

First order hidden Markov models - Lerning problem



First order hidden Markov models

- Context prediction with hidden Markov models:
 - Given the transition model, estimate all p_{ij} and b_{jk}
 - Given a sequence V^T, decode the most probable sequence of hidden states
 - Sector 2 Sec

Outline

Markov prediction approaches

- Introduction and Markov properties
- 2 Markov chains
- O Hidden Markov Models
- Ontext prediction with Markov approaches
 - Properties of Markov prediction approaches
- Conditional random fields
 - Context prediction with CRF
 - Properties of CRF prediction approaches

6 Conclusion

Context prediction with Markov approaches

- Given: Sequence of contexts $\xi_{0-k+1}, \ldots, \xi_0$
- Generate Markov chain representing the transition probabilities for each pair of observations
- Now possible: Provide probability distribution on the next outcome
- Can also be generalised to higher order Markov processes
- Several iterations of this process provide higher prediction horizons

Memory and processing load

- Runtime dependent on size of probability graph G
- C: Set of different context values
- The number of states of the Markov chain: *C*.
- Time to find most probable next state is O(|C|) in the worst case.
 - Every arc to possible following context to be considered.
 - |C| 1 arcs existent in the worst case.
- Most probable *n* future context time series elements
 - Naive computation time: $O(|C|^n)$
 - When transition probabilities stored to a matrix, one matrix multiplication for every future context
 - Computation time: $O(n \cdot |C|^2)$.

Memory and processing load

- Memory requirements
 - Dependent on the number of contexts observed size of the transition matrix
 - Order 1: $O(|C|^2)$
 - Order k: $O(|C|^{k+1})$

Prediction horizon

- Prediction horizon can be extended by iterative prediction
 - Utilise predicted contexts as input
- Problem: Less accurate
 - Predicted contexts more error prone than measured values

Properties of Markov prediction approaches Adaptability

- The Markov prediction approach is well able to adapt to changing environments
 - Adapt context transition probabilities
 - Consideration of new events possible
 - Rebuild of transition matrix required

Multi-dimensional time series

- The Markov prediction algorithm is not suited for multi-dimensional time series
 - Designed for one-dimensional Input
 - Possible: Aggregation of multi-dimensional time series to one-dimensional time series.

Iterative prediction

• Iterative Prediction possible

- Steep decrease in prediction accuracy expected since prediction horizon is only 1
- Increase of prediction horizon possible by Aggregation of context sequence of fixed length in one Markov state
 - Prediction horizon fixed
 - Increase in Memory consumption and processing time
 - When I contexts are aggregated: I^C states
 - Runtime:

$$O(n \cdot l^{C^2}).$$

• Memory consumption:

```
O(I^{C^2}) (order one)
O(I^{C^{k+1}}) (order k)
```

Prediction of context durations

- Prediction of context duration not possible
 - Only simple sequence of occurring contexts possible

Approximate matching of patterns

• Exact pattern matching

• The Markov prediction algorithm utilises exact pattern matching

Context data types

- All context data types supported
 - Every distinct context type one state
 - Probably drastic increase in runtime and memory consumption for numeric context types
 - Possible: Assign intervals to states

Pre-processing

- Pre-processing required to construct context transition probabilities
- On-line approach feasible learning
- Runtime: O(k)
 - Count frequency of specific context transitions in training time series of length \boldsymbol{k}

Aspects of prediction algorithms

Summary

	IPAM	ONISI	Markov	CRF
Numeric Contexts	yes	no	yes	
Non-numeric Contexts	yes	yes	yes	
Complexity	O(k)	()	$O(C^2)$	
Learning ability	(no)	yes	yes	
Approximate matching	no	no	no	
Multi-dim. TS	(no)	(no)	(no)	
Discrete data	yes	yes	yes	
Variable length patterns	no	yes	no	
Multi-type TS	yes	no	(no)	
Continuous data	no	no	no	
Pre-processing	O(k)	-	O(k)	
Context durations	no	no	no	
Continuous time	no	no	yes	

Algorithms for context prediction in Ubiquitous Systems

- Markov processes are straightforward and easily applied to context prediction tasks.
- Model can be applied to numerical and non-numerical data alike.
- Prediction that reaches farther into future implicitly utilises already predicted data which might consequently decrease the prediction accuracy.

Outline

Markov prediction approaches

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6 Conclusion

Introduction

- <u>Undirected</u> graphical model ^{6 7}
- Similar to HMM
 - HMM specific CRF
 - Relax assumptions about input and output sequence
 - Instead of constant transition probability: Arbitrary functions that vary across positions in sequence of hidden states
- Vertices represent random variables
- Edges represent dependency between two random variables

⁶ John Lafferty, Andrew McCallum and Fernando Pereira, *Conditional random fields: Probabilistic models for segmenting and labeling sequence data*, In Proceedings of the 18th international conference on machine learning, pp 282-289, 2001.

⁷Douglas L. Vail, Manuela M. Veloso and John D. Lafferty, *Conditional random fields for activity recognition*, In Proceedings of the AAMAS, 2007.

Introduction

- Layout of inner-graph (hidden states) arbitrary
- Input sequence: X
- Inner states: Y
- Conditional dependency of each Y_i on X defined through set of feature functions

$$f(i, Y_{i-1}, Y_i, X)$$
 (36)

• Each feature assigned numerical weight

Algorithms for context prediction in Ubiquitous Systems

Training

- Various learning approaches to train CRF
 - Gradient based
 - Quasi-Newton-approach
- Sequences provided of which also desired output is known
- CRF parameters are adapted to match a maximum number of training sequences

Applications

- Applications similar to HMM
- Labeling or parsing of sequential data
 - Natural language text
 - Biological sequences
 - Classification of proteins
 - Prediction of the secondary structure of DNS and proteine
 - Image recognition and image resauration

Discussion

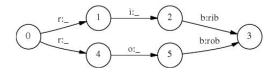
- Generative models
 - HMMs, stochastic grammars, ...
 - Assign joint probability to paired observations
- Discriminative models
 - Maximum entropy Markov models, Conditional random fields,

• •

Discussion

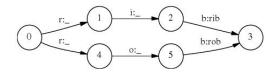
• Problem of generative models

- To define joint probability over observation sequences (e.g. words or nucleotides), all possible observation sequences are enumerated
- Not practical
 - Multiple interacting features
 - Long range dependencies
- Conditional probability depends on fixed, dependent features
 - Instead of arbitrary independent features
- Very strict independence assumptions on observations
 - e.g. conditional independence

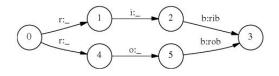


- Problem of (classical) discriminative models (e.g. MEMM)
 - Label bias problem⁸
 - Conservation of score mass
 - States with fewer outgoing transitions are tendentially biased
 - States with low-entropy next state distributions will take little notice of observations

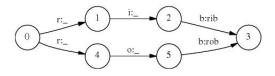
⁸L. Bottou, Une approche theorique de l'apprentissage connexionniste: Applications a la reconnaissance de la parole, PhD-thesis, 1991.



- Solutions proposed to solve the label bias problem
 - Change state transition structure of the model
 - Collapse states (here: 1 and 4)
 - Delay branching until discriminating observation
 - Start with fully connected model
 - Let training figure out good structure



- Problems with the solutions proposed
 - Change state transition structure of the model
 - Not always possible
 - May lead to combinatorical explosion
 - Start with fully connected model
 - Preludes use of valuable prior structural knowledge



- Requirement for proper solutions
 - Model that accounts for whole state sequences at once
 - Let transitions 'vote' more strongly than others
 - Score mass will not be conserved
 - Transitions can amplify or dampen received mass

Discussion

- Conditional random fields
 - Solve label bias problem
 - Single exponential model for joint probability of sequences
 - Less impacted by higher-order dependencies between states

Algorithmic model

Definition: Conditional random field

Let G = (V, E) be a graph such that $Y = (Y_v)_{v \in V}$. Then (X, Y) is a conditional random field when the random variables Y_v obey the Markov property with respect to the graph: $p(Y_v|X, Y_w, w \neq v) = p(Y_v|X, Y_w, w \sim v)$

- $w \sim v$ means that w and v are neighbours in G
- A CRF is a random field conditioned on the observation X
- In general, the graphical structures of X and Y are not the same.

Random field

- Random field
 - Generalization of a stochastic process
 - Underlying parameter need no longer be a simple real
 - Can instead be multidimensional vector space

Random field

Random field

Let $S = X_1, ..., X_n$, with the X_i in $\{0, 1, ..., G - 1\}$ being a set of random variables on the sample space $\Omega = \{0, 1, ..., G - 1\}^n$. A probability measure π is a random field if, for all ω in Ω , $\pi(\omega) > 0$.

Example: HMM-like CRF

$$f_{y',y}(\langle u, v \rangle, y |_{\langle u, v \rangle}, x) = \delta(y_u, y')\delta(y_v, y)$$
(37)
$$g_{y,x}(v, y |_v, x) = \delta(y_v, y)\delta(x_v, x)$$
(38)

• Feature for each state pair (y, y') and each state-observation pair (y, x)

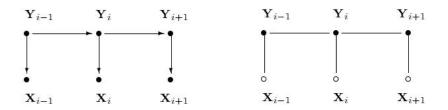
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Example: HMM-like CRF

- CRF more expressive: Can model more cases than HMM
- Features do not need to specify completely a state or observation.
 - Therefore, model can be estimated from less training data
- CRFs share all convexity properties of general maximum entropy models

Example: HMM-like CRF



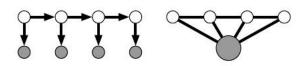
• Graphical structures of HMMs and CRFs

• Circle indicates that variable is not generated by the model

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Example: HMM-like CRF



• CRF: Entire observation sequence combined

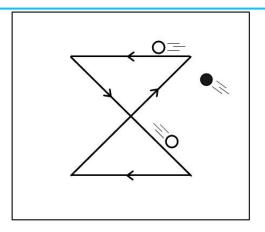
Experiments

model	error	oov error
HMM	5.69%	45.99%
MEMM	6.37%	54.61%
CRF	5.55%	48.05%
MEMM ⁺	4.81%	26.99%
CRF ⁺	4.27%	23.76%

⁺Using spelling features

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Experiments



- Seeker robot (black) tries to tag on of the other players
- Simplified variant: Non-seeker robots follow hourclass pattern

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Experiments

• CRF and HMM accuracy for identifying the seeker

		Hourglass		Unconstrained			
Features	HMM Acc.	CRF Acc.	$\ell(Y X)$	HMM Acc.	CRF Acc.	$\ell(Y X)$	
Positions	33.1	53.6	-959.7	37.1	37.8	-1354.5	
Velocities	68.4	89.4	-717.1	55.7	70.4	-1206.5	
Velocity Thresholds							
$W = \frac{1}{60}$ th sec.	62.5	71.2	-818.0	46.8	58.6	-1148.6	
W = 0.1 sec.	63.0	73.9	-784.3	46.0	62.4	-1099.2	
W = 0.5 sec.	63.6	80.6	-708.8	68.9	71.9	-983.1	
W = 1.0 sec.	60.2	83.1	-721.8	67.8	75.3	-980.9	
W = 1.5 sec.	56.9	85.5	-731.7	68.8	77.8	-1004.7	
W = 2.0 sec.	53.7	87.1	-751.1	72.1	77.3	-1036.3	
Chasing	75.8	95.4	-622.3	65.5	80.4	-1058.3	
Distance (U)	46.6	49.5	-779.7	43.5	42.3	-604.4	
Distance (N)	46.6	49.9	-200.6	43.5	58.0	-223.4	
Distance & Chasing (U)	75.6	99.3	-90.6	65.8	93.9	-181.8	
Distance & Chasing (N)	75.6	99.3	-115.3	65.8	97.6	-112.2	
Many Features	72.4	98.1	-172.2	63.4	98.5	-178.9	
Redundant Features	72.4	95.7	-509.3	52.7	93.8	-6432.3	

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Algorithms for context prediction in Ubiquitous Systems

Context prediction with CRF

Prediction procedure

- Context prediction with CRF:
 - Given the transition model, estimate all probabilites between states and state action probabilites
 - Given a sequence V^T, decode the most probable sequence of hidden states
 - Sector 2 Sec

Aspects of prediction algorithms

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Continuous data	no	no	no	no
Pre-processing	O(k)	_	O(k)	O(k)
Context durations	no	no	no	no
Continuous time	no	no	yes	yes

Algorithms for context prediction in Ubiquitous Systems

Properties of CRF prediction approaches

Conclusion

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Conclusion