# Algorithms for context prediction in Ubiquitous Systems 

Markov prediction approaches

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## Overview and Structure

- Introduction to context aware computing
- Basics of probability theory
- Algorithms
- Simple prediction approaches: ONISI and IPAM
- Markov prediction approaches
- The State predictor
- Alignment prediction
- Prediction with self organising maps
- Stochastic prediction approaches: ARMA and Kalman filter
- Alternative prediction approahces
- Dempster shafer
- Evolutionary algorithms
- Neural networks
- Simulated annealing


## Overview and Structure

- Introduction to context aware computing
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## Outline

## Markov prediction approaches

(1) Introduction and Markov properties
(2) Markov chains
(3) Hidden Markov Models

4 Context prediction with Markov approaches

- Properties of Markov prediction approaches
(5) Conditional random fields
- Context prediction with CRF
- Properties of CRF prediction approaches
(6) Conclusion


## Introduction and Markov properties

## Historical remarks

- Markov processes
- Intensively studied
- Major branch in the theory of stochastic processes
- A. A. Markov (1856-1922)
- Extended by A. Kolmogorov by chains of infinitely many states
- 'Anfangsgründe der Theorie der Markoffschen Ketten mit unendlich vielen möglichen Zuständen' (1936) ${ }^{1}$

[^0]
## Introduction and Markov properties

## Historical remarks

- Markov Chains
- Theory of Markov chains applied to a variety of algorithmic problems
- Standard tool in many probabilistic applications
- Intuitive graphical representation
- Possible to illustrate properties of stochastic processes graphically
- Popular for their simplicity and easy applicability to huge set of problems ${ }^{2}$

[^1]
## Introduction and Markov properties

## Introcution

- Dependent trials of events
- Set of possible outcomes of a measurement $E_{i}$ associated with occurence probability $p_{i}$
- When occurence of events is not independent
- Probability to observe specific sequence $E_{1}, E_{2}, \ldots, E_{i}$ obtained by conditional probability:

$$
\begin{equation*}
P\left(E_{i} \mid E_{1}, E_{2}, \ldots, E_{i-1}\right) \tag{1}
\end{equation*}
$$

- In general:

$$
\begin{equation*}
P\left(E_{i} \mid E_{1}, E_{2}, \ldots, E_{i-1}\right) \neq P\left(E_{i} \mid E_{2}, \ldots, E_{i-1}\right) \tag{2}
\end{equation*}
$$

## Introduction and Markov properties

## Independent random variables

- Sequence of tials for independent random variable
- $T$ : number of trials up to first success of probability $p$.
- Then:

$$
\begin{equation*}
P\{T>k\}=(1-p)^{k} \tag{3}
\end{equation*}
$$

- Suppose: No success during the first $m$ trials
- Waiting time $T$ to first success for $m$-th trial has same distribution $(1-p)^{k}$
- Independent of number of preceding failures $m$


## Introduction and Markov properties

## Examples

- Independent random variables
- Number of coin tosses until 'head' is observed
- Radioactive atoms always have the same probability of decaying at the next trial
- Dependent random variables
- The knowledge that no streetcar has passed for five minutes increases our expectation that it will come soon.
- Coin tossing:
- Probability that the cumulative numbers of heads and tails will equalize at the second trial is $\frac{1}{2}$
- Given that they did not, the probability that they equalize after two additional trials is only $\frac{1}{4}$


## Introduction and Markov properties

## Lack of memory - Rigorous formulation

- Suppose a waiting time $T$ assumes the values $0,1,2, \ldots$ with probabilities $p_{0}, p_{1}, p_{2}, \ldots$
- Let $T$ have the following property
- Conditional probability that the waiting time terminates at the $k$-th trial equals $p_{0}$
- Then:
- $p_{k}=\left(1-p_{0}\right)^{k} p_{0}$


## Introduction and Markov properties

## Lack of Memory - Rigorous formulation

## Proof.

- $1-p_{k}=p_{k+1}+p_{k+2}+\cdots=P\{T>k\}$
- Conditional probability of $T=k: p_{k} /\left(1-p_{k-1}\right)$
- Assumption for all $k \geq 1: \frac{p_{k}}{1-p_{k-1}}=p_{0}$
- Since $p_{k}=\left(1-p_{k-1}\right)-\left(1-p_{k}\right)$

$$
\begin{equation*}
\frac{1-p_{k}}{1-p_{k-1}}=1-p_{0} \tag{4}
\end{equation*}
$$

- since $1-p_{0}=p_{1}+p_{2}+\ldots: 1-p_{k}=\left(1-p_{0}\right)^{k+1}$ and

$$
\begin{equation*}
p_{k}=1-p_{k-1}-\left(1-p_{k}\right)=\left(1-p_{0}\right)^{k} p_{0} \tag{5}
\end{equation*}
$$

## Introduction and Markov properties

## Markov property

Markov property
In the theory of stochasitc processes the described lack of memory is connected with the Markovian property.

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## Markov chains

## Dependence and independence of events

- Independent trials of events
- Set of possible outcomes of a measurement $E_{i}$ associated with occurence probability $p_{i}$
- Probability to observe sample sequence:
- $P\left\{\left(E_{1}, E_{2}, \ldots, E_{i}\right)\right\}=p_{1} p_{2} \cdots p_{i}$


## Markov chains

## Dependence and independence of events

- Theory of Markov chains:
- Outcome of any trial depends exclusively on the outcome of the directly preceding trial
- Outcome of $E_{k}$ is no longer associated with fixed probability $p_{k}$
- Instead: With every pair $\left(E_{i}, E_{j}\right)$ a conditional probability $p_{i j}$
- Probability that $E_{j}$ is observed after $E_{i}$
- Additionally: Probability $a_{i}$ of the event $E_{i}$


## Markov chains

## Dependence and independence of events

- Theory of Markov chains:
- $P\left\{\left(E_{i}, E_{j}\right)\right\}=a_{i} p_{i j}$
- $P\left\{\left(E_{i}, E_{j}, E_{k}\right)\right\}=a_{i} p_{i j} p_{j k}$
- $P\left\{\left(E_{i}, E_{j}, E_{k}, E_{l}\right)\right\}=a_{i} p_{i j} p_{j k} p_{k l}$
- $P\left\{\left(E_{i}, E_{j}, \ldots, E_{m}, E_{n}\right)\right\}=a_{i} p_{i j} \ldots p_{m n}$


## Markov chains

## Markov chain

## Markov chain

A sequence of observations $E_{1}, E_{2}, \ldots$ is called a Markov chain if the probabilities of sample sequences are defined by

$$
\begin{equation*}
P\left(E_{1}, E_{2}, \ldots, E_{i}\right)=a_{1} \cdot p_{12} \cdot p_{23} \cdots \cdot p_{(i-1) i} \tag{6}
\end{equation*}
$$

and fixed conditional probabilities $p_{i j}$ that the event $E_{i}$ is observed directly in advance of $E_{j}$.

## Markov chains

## Markov chain

- Markov chain described by probability a for initial distribution and matrix $P$ of transition probabilities.

$$
P=\left[\begin{array}{cccc}
p_{11} & p_{12} & p_{13} & \cdots  \tag{7}\\
p_{21} & p_{22} & p_{23} & \cdots \\
\vdots & \vdots & \vdots & \ddots
\end{array}\right]
$$

- $P$ is a square matrix with non-negative entries that sum to 1 in each row.


## Markov chains

## Stochastic matrix

- $P$ is called a stochastic matrix.
- Any stochastic matrix is suited to describe transition probabilities of Markov chains.


## Markov chains

## Markov chain

- Markov chain sometimes modelled as directed graph $G=(V, E)$
- Labelled edges in $E$
- states (or vertices) in $V$.
- Transition probabilities $p_{i j}$ between $E_{i}, E_{j} \in V$



## Markov chains

## Derive state transision probabilities

- $p_{i j}^{k}$ denotes probability that $E_{j}$ is observed exactly $k$ observations after $E_{i}$ was observed.
- Calculated as the sum of the probabilities for all possible paths $E_{i} E_{i_{1}} \cdots E_{i_{k-1}} E_{j}$ of length $k$
- We already know

$$
\begin{equation*}
p_{i j}^{1}=p_{i j} \tag{8}
\end{equation*}
$$

- Consequently:

$$
\begin{equation*}
P_{i j}^{2}=\sum_{\nu} p_{i \nu} \cdot p_{\nu j} \tag{9}
\end{equation*}
$$

## Markov chains

## Derive state transision probabilities

- By mathematical induction:

$$
\begin{equation*}
p_{i j}^{n+1}=\sum_{\nu} p_{i \nu} \cdot p_{\nu j}^{n} \tag{10}
\end{equation*}
$$

- and

$$
\begin{equation*}
p_{i j}^{n+m}=\sum_{\nu} p_{i \nu}^{m} \cdot p_{\nu j}^{n}=\sum_{\nu} p_{i \nu}^{n} \cdot p_{\nu j}^{m} \tag{11}
\end{equation*}
$$

## Markov chains

## Derive state transision probabilities

- Similar to the matrix $P$ we can create a matrix $P^{n}$ that contains all $p_{i j}^{n}$
- We obtain $P_{i j}^{n+1}$ from $P^{n+1}$ by multiplying all elements of the $i$-th row of $P$ with the correspoinding elements of the $j$-the column of $P^{n}$ and add all products.
- Symbolically: $P^{n+m}=P^{n} P^{m}$.


## Markov chains

## Examples

- Markov chains:
- Urn models
- Every Markov chain is equivalent to an urn model
- Each urn represents a state in a markov chain and probabilities to draw specific balls represent possible events in this state
- Branching processes
- Instead of saying that the $n$-th trial results in $E_{k}$ we say that the $n$-th generation is of size $k$
- Random walk on a line
- Events are transitions between states
- Only two events are possible in each state


## Markov chains

## Random walks and ruin problems

- Random walk
- When there are only two possible states $E_{1}$ and $E_{2}$ the matrix of transition probabilities is of the form

$$
P=\left[\begin{array}{cc}
1-p & p  \tag{12}\\
\alpha & 1-\alpha
\end{array}\right]
$$

- Can be realised by particle moving along one axis in one or the other direction.
- System is in state $E_{1}$ when the particle moves into one direction and in state $E_{2}$ otherwise.


## Markov chains

## Random walks and ruin problems

- Possible problems / questions
- Expected time to return to origin
- Expected time to return to origin given that the starting point had a specific distance to the origin
- ...


## Markov chains

## Random walks and ruin problems

- Random walk with absorbing barriers

$$
P=\left[\begin{array}{cccccccc}
1 & 0 & 0 & 0 & \cdots & 0 & 0 & 0  \tag{13}\\
1-p & 0 & p & 0 & \cdots & 0 & 0 & 0 \\
0 & 1-p & 0 & p & \cdots & 0 & 0 & 0 \\
& & & & \vdots & & & \\
0 & 0 & 0 & 0 & \cdots & 1-p & 0 & p \\
0 & 0 & 0 & 0 & \cdots & 0 & 0 & 1
\end{array}\right]
$$

- First and last state are obsorbing
- All inner states implement a random walk on the line
- Possible application: Game between two players with equal money balance where the loosing one has to pay one unit to the winner.


## Markov chains

## Random walks and ruin problems

- Random walk with reflecting barriers

$$
P=\left[\begin{array}{cccccccc}
1-p & p & 0 & 0 & \cdots & 0 & 0 & 0  \tag{14}\\
1-p & 0 & p & 0 & \cdots & 0 & 0 & 0 \\
0 & 1-p & 0 & p & \cdots & 0 & 0 & 0 \\
& & & & \vdots & & & \\
0 & 0 & 0 & 0 & \cdots & 1-p & 0 & p \\
0 & 0 & 0 & 0 & \cdots & 0 & 1-p & p
\end{array}\right]
$$

- First and last state are reflecting
- All inner states implement a random walk on the line


## Markov chains

## Random walks and ruin problems

## Classical ruin problem

- Consider a gambler who wins or loses a dollar with probabilities $p$ and $1-p$
- Initial capital of gambler and adversary: $z, a-z$
- Game ends when the capital reaches 0 or $a$.
- When one of the players is ruined
- We are interested in the probability of the gamblers ruin and the probability distribution of the game


## Markov chains

## Random walks and ruin problems

- Gamblers ruin problem
- Random walk with absorbing barriers at 0 and $a$
- Examples:
- Physicists use this model as crude approximation to one-dimensional diffusion or Brownian motion (Particle is exposed to great number of molecular collisions which impart to it a random motion)
- $p>1 / 2$ represents a drift to the right, when shocks from the left are more probable


## Markov chains

## Random walks and ruin problems

- Probability of gamblers ruin
- $q_{z}$ : Probability of gambler's ultimate ruin when $z$ is the starting capital and $a$ is the overal capital
- After the first trial the gablers's fortune is either $z-1$ or $z+1$ :

$$
q_{z}=p q_{z+1}+(1-p) q_{z-1}
$$

- We can show:

$$
\begin{align*}
& p \neq \frac{1}{2}: \quad q_{z}=\frac{\left(\frac{1-p}{p}\right)^{a}-\left(\frac{1-p}{p}\right)^{z}}{\left(\frac{1-p}{p}\right)^{a}-1}  \tag{15}\\
& p=\frac{1}{2}: \quad q_{z}=1-\frac{z}{a} \tag{16}
\end{align*}
$$

## Markov chains

## Random walks and ruin problems

- Probability of gamblers ruin
- The probability $p_{z}$ of the gambler winning the game is equal to the probability of his adversary loosing the game.
- It is therefore obtained in the same way by replacing $p$ with $1-p$ and $z$ by $a-z$
- Therefore: $p_{z}+q_{z}=1$


## Markov chains

## Random walks and ruin problems

- Some interesting results
- Since for $p=\frac{1}{2}$, we have derived $q_{z}=1-\frac{z}{a}$
- A player with initial capital $z=999$ has a probability of 0.999 to win a dollar before losing his capital.
- With $p=0.4$ the game is unfavorable, but still the probability of winning a dollar before losing the capital is about $\frac{2}{3}$


## Markov chains

## Random walks and ruin problems

## Example - anecdote

A certain man used to visit Monte Carlo year after year and was always successful in recovering the cost of his vacations. He firmly believed in a magic power over chance.

This experience is not surprising.

- Assuming that he started with ten times the ultimate gain, the chances of success in any year are nearly 0.9.
- The probability of an unbroken sequence of ten successes is about $\left(1-\frac{1}{10}\right)^{10} \approx e^{-1} \approx 0.37$
Therefore, continued success is by no means improbable
- However: one failue would result in the gambler's ruin :-)


## Markov chains

## Random walks and ruin problems

lllustrating the Classical Ruin Problem

| $p$ | $q$ | $z$ | $a$ | Probability of |  | Expected |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Ruin | Success | Gain | Duration |
| 0.5 | 0.5 | 9 | 10 | 0.1 | 0.9 | 0 | 9 |
| 0.5 | 0.5 | 90 | 100 | 0.1 | 0.9 | 0 | 900 |
| 0.5 | 0.5 | 900 | 1,000 | 0.1 | 0.9 | 0 | 90,000 |
| 0.5 | 0.5 | 950 | 1,000 | 0.05 | 0.95 | 0 | 47,500 |
| 0.5 | 0.5 | 8,000 | 10,000 | 0.2 | 0.8 | 0 | 16,000,000 |
| 0.45 | 0.55 | 9 | 10 | 0.210 | 0.790 | $-1.1$ | 11 |
| 0.45 | 0.55 | 90 | 100 | 0.866 | 0.134 | -76.6 | 765.6 |
| 0.45 | 0.55 | 99 | 100 | 0.182 | 0.818 | -17.2 | 171.8 |
| 0.4 | 0.6 | 90 | 100 | 0.983 | 0.017 | -88.3 | 441.3 |
| 0.4 | 0.6 | 99 | 100 | 0.333 | 0.667 | $-32.3$ | 161.7 |
| The initial capital is $z$. The game terminates with ruin (loss z) or capital $a$ (gain $a-z$ ) |  |  |  |  |  |  |  |

- Effect of increasing stakes is more pronounced than might be expected


## Markov chains

## Random walks and ruin problems

- Expected duration of the game
- $D_{z}$ : Expected duration of the game when $z$ is the starting capital and $a$ is the overal capital
- After the first trial the gablers's fortune is either $z-1$ or $z+1$ :

$$
\text { - } D_{z}=p D_{z+1}+(1-p) D_{z-1}+1
$$

- We can show:

$$
\begin{align*}
& p \neq \frac{1}{2} \quad: \quad D_{z}=\frac{z}{1-2 p}-\frac{a}{1-2 p} \cdot \frac{1-\left(\frac{1-p}{p}\right)^{z}}{1-\left(\frac{1-p}{p}\right)^{a}}  \tag{17}\\
& p=\frac{1}{2} \quad: \quad D_{z}=z(a-z) \tag{18}
\end{align*}
$$

## Markov chains

## Random walks and ruin problems

- Expected duration of the game

$$
\begin{align*}
& p \neq \frac{1}{2} \quad: \quad D_{z}=\frac{z}{1-2 p}-\frac{a}{1-2 p} \cdot \frac{1-\left(\frac{1-p}{p}\right)^{z}}{1-\left(\frac{1-p}{p}\right)^{a}}  \tag{19}\\
& p=\frac{1}{2} \quad: \quad D_{z}=z(a-z) \tag{20}
\end{align*}
$$

Examples - Duration considerably longer as naively expected:

- If two players with 500 dollars each toss a fair coin, average duration of the game is 250000 trials
- If a gambler has only one dollar and his adversary 1000, the average duration is 1000 trials


## Markov chains

Closures and closed sets

## Closed set of states

- A set $C$ of states is closed if no state outside $C$ can be reached from any state $E_{i}$ in $C$.
- For an arbitrary set $C$ of states the smallest closest set containing $C$ is called the closure of $C$
- A single state $E_{k}$ forming a closed set is called absorbing


## Markov chains

Closures and closed sets

Closed sets in stochastic matrices
If in a matrix $P^{n}$ all rows and all columns corresponding to states outside a closed set $C$ are deleted, the remaining matrices are again stochastic matrices.

## Markov chains

Closures and closed sets

Irreducible Markov chain
A Markov chain is irreducible if there exists no closed set other than the set of all states.

Criterion for irreducible chains
A chain is irreducible if, and only if, every state can be reached from every other state.

## Markov chains

## Periodicity

Periodicity of states

- The state $E_{j}$ has period $t>1$ if $p_{j j}^{n}=0$ unless $n=v t$ is a multiple of $t$ and $t$ is the largest integer with this property.
- the state $E_{j}$ is aperiodic if no such $t>1$ exists
- A state $E_{j}$ to which no return i possible is considered aperiodic


## Markov chains

## Periodicity

- To deal with a periodic $E_{j}$ it suffices to consider the chain at the trials $t, 2 t, 3 t$
- In this way we obtain a new Markov chain with transition probabilities $p_{i k}^{t}$
- In this new chain $E_{j}$ is aperiodic
- Results concerning aperiodic states can thus be transferred to periodic states


## Markov chains

## Persistent and transient states

## Persistent and transient states

- The state $E_{j}$ is persistent if $\sum_{n=1}^{\infty} p_{j j}^{n}=1$ and transient if $\sum_{n=1}^{\infty} p_{j j}^{n}<1$
- A persistent state $E_{j}$ is called null state if its mean recurrrence time $\mu_{j}=\infty$


## Markov chains

## Irreducible chains

- Two states are of the same type when they are either
- both aperiodic
- both have the same period
- both are transient
- both are persistent and each
- with infinite recurrence times
- or finite recurrence times


## Markov chains

## Irreducible chains

Type of states in irreducible chains
All states of an irreducible chain are of the same type

## Markov chains

## Applications - Card shuffling

- A deck of $N$ cards can be arranged in $N$ ! different orders.
- Each order represent a possible state of the system
- We conceive each particular shuffling operation as a transformation $E_{i} \rightarrow E_{j}$
- Result:
- The permutation is not cyclic
- Therefore, repeated application of a single operation will never visit all possible states
- This means that the original state is again observed before all states are visited
- This is a Markov chain:
- We assume that a player applies several shuffling operation with a random probability and that the current order of the cards is not known.


## Markov chains

Markov processes

## Markov process

A sequence of discrete-valued random variables is a markov process if the joint distribution of $\left(X^{1}, \ldots, X^{n}\right)$ is defined in such a way that the conditional probability of the relation $X^{n}=x$ on the hypothesis $X^{n_{1}}=x_{1}, \ldots, X^{n_{r}}=x_{r}$ is identical with the conditional probability of $X^{n}=x$ on the single hypothesis $X^{n_{r}}=x_{r}$.

## Markov chains

## Higher order Markov processes

- Order k Markov processes
- Typically
- Occurence of event dependent on $k$ events that were observed directly beforehand
- Constrained lack of memory
- Dependence between the last $k$ events observed
- Useful for context prediction / time series forecasting, when typical patterns or trends are to be considered


## Markov chains

## Higher order Markov processes

- Probability that $E_{1}, E_{2}, \ldots, E_{i}$ observed is then

$$
\begin{equation*}
P\left(E_{1}, E_{2}, \ldots, E_{i}\right)=p_{1} \cdot p_{12} \cdot p_{23} \cdots \cdot p_{(i-1) i} \tag{21}
\end{equation*}
$$

- Required: $p_{i}>0 \forall i$ and $\sum_{p_{i}}=1$.


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## Hidden Markov Models

- Make a sequence of decisions for a process that is not directly observable ${ }^{3}$
- Current states of the process might be impacted by prior states
- HMM often utilised in speach recognition or gesture recognition

[^2]
## Hidden Markov Models

## Applications

- Computational biology
- Align biological sequences
- Find sequences homologous to a known evolutionary family
- Analyse RNA secondary structure ${ }^{4}$
- Computational linguistics ${ }^{5}$
- Topic segmentation of text
- Information extraction

[^3]
## Hidden Markov Models

First order Markov models


## Hidden Markov Models

First order hidden Markov models



## Hidden Markov Models

## First order hidden Markov models



- At every time step $t$ the system is in an internal state $\omega(t)$
- Additionally, we assume that it emits a (visible) symbol $v(t)$
- Only access to visible symbols and not to intenal states


## Hidden Markov Models

## First order hidden Markov models

- $V^{T}=\{v(1), v(2), \ldots, v(T)\}$
- In any state $\omega(t)$ we have a probability of emitting a particular visible symbol $v_{k}(t)$
- Probability to be in state $\omega_{j}(t)$ and emit symbol $v_{k}(t)$ :
- $P\left(v_{k}(t) \mid \omega_{j}(t)\right)=b_{j k}$
- Transmission probabilities: $p_{i j}=P\left(\omega_{j}(t+1) \mid \omega_{i}(t)\right)$
- Emission probability: $b_{j k}=P\left(v_{k}(t) \mid \omega_{j}(t)\right)$


## Hidden Markov Models

## First order hidden Markov models

- Normalisation conditions

$$
\begin{align*}
\sum_{j} p_{i j} & =1 \text { for all } i  \tag{22}\\
\sum_{k} b_{j k} & =1 \text { for all } j \tag{23}
\end{align*}
$$

## Hidden Markov Models

## First order hidden Markov models

- Central issues in hidden Markov models:
- Evaluation problem
- Determine the probability that a particular sequence of visible symbols $V^{T}$ was generated by a given hidden Markov model
- Decoding problem
- Determine the most likely sequence of hidden states $\omega^{T}$ that led to a specific sequence of observations $V^{T}$
- Learning problem
- Given a set of training observations of visible symbols, determine the parameters $p_{i j}$ and $b_{j k}$ for a given HMM


## Hidden Markov Models

First order hidden Markov models - Evaluation problem

- Probability that model produces a sequence $V^{T}$ :

$$
\begin{equation*}
P\left(V^{T}\right)=\sum_{\bar{\omega}^{T}} P\left(V^{T} \mid \bar{\omega}^{T}\right) P\left(\bar{\omega}^{T}\right) \tag{24}
\end{equation*}
$$

- Also:

$$
\begin{align*}
P\left(\bar{\omega}^{T}\right) & =\prod_{t=1}^{T} P(j \omega(t) \mid \omega(t-1))  \tag{25}\\
P\left(V^{T} \mid \bar{\omega}^{T}\right) & =\prod_{t=1}^{T} P(v(t) \mid \omega(t)) \tag{26}
\end{align*}
$$

- Together:

$$
\begin{equation*}
P\left(V^{T}\right)=\sum_{\bar{\omega}^{T}} \prod_{t=1}^{T} P(v(t) \mid \omega(t)) P(\omega(t) \mid \omega(t-1)) \tag{27}
\end{equation*}
$$

## Hidden Markov Models

First order hidden Markov models - Evaluation problem

- Probability that model produces a sequence $V^{T}$ :

$$
\begin{equation*}
P\left(V^{T}\right)=\sum_{\bar{\omega}^{T}} \prod_{t=1}^{T} P(v(t) \mid \omega(t)) P(\omega(t) \mid \omega(t-1)) \tag{28}
\end{equation*}
$$

- Formally complex but straightforward
- Naive computational complexity
- $O\left(c^{T} T\right)$


## Hidden Markov Models

## First order hidden Markov models - Evaluation problem

- Probability that model produces a sequence $V^{T}$ :

$$
\begin{equation*}
P\left(V^{T}\right)=\sum_{\bar{\omega}^{T}} \prod_{t=1}^{T} P(v(t) \mid \omega(t)) P(\omega(t) \mid \omega(t-1)) \tag{29}
\end{equation*}
$$

- Computationally less complex algorithm:
- Calculate $P\left(V^{T}\right)$ recursively
- $P(v(t) \mid \omega(t)) P(\omega(t) \mid \omega(t-1))$ involves only $v(t), \omega(t)$ and $\omega(t-1)$

$$
\alpha_{j}(t)= \begin{cases}0 & t=0 \text { and } j \neq \text { initial state }  \tag{30}\\ 1 & t=0 \text { and } j=\text { initial state } \\ {\left[\sum_{i} \alpha_{i}(t-1) p_{i j}\right] b_{j k} v(t)} & \text { otherwise }\end{cases}
$$

## Hidden Markov Models

First order hidden Markov models - Evaluation problem

- Forward Algorithm
- Computational complexity: $O\left(c^{2} T\right)$


## Forward algorithm

```
1 initialise t}\leftarrow0,\mp@subsup{p}{ij}{},\mp@subsup{b}{jk}{},\mp@subsup{V}{}{T},\mp@subsup{\alpha}{j}{}(o
2 for t \leftarrowt+1
3 < <j(t)\leftarrow bjk v(t) \sum i=1
4 \mp@code { u n t i l ~ t = T }
5 \text { return } P ( V ^ { T } ) \leftarrow \alpha _ { 0 } ( T ) \text { for the final state}
6 end
```


## Hidden Markov Models

First order hidden Markov models - Decoding problem

- Given a sequence $V^{T}$, find the most probable sequence of hidden states.
- Enumeration of every possible path will cost $O\left(c^{T} T\right)$
- Not feasible


## Hidden Markov Models

First order hidden Markov models - Decoding problem

- Given a sequence $V^{T}$, find the most probable sequence of hidden states.

```
Decoding algorithm
1 initialise path \(\leftarrow\}, t \leftarrow 0\)
2 for \(t \leftarrow t+1\)
\(3 \quad j \leftarrow j+1\)
\(4 \quad\) for \(j \leftarrow j+1\)
\(5 \quad \alpha_{j}(t) \leftarrow b_{j k} v(t) \sum_{i=1}^{c} \alpha_{i}(t-1) p_{i j}\)
\(6 \quad\) until \(j=c\)
\(7 \quad j^{\prime} \leftarrow \arg \max _{j} \alpha_{j}(t)\)
8 append \(\omega_{j^{\prime}}\) to path
9 until \(t=T\)
10 return path
11 end
```


## Hidden Markov Models

## First order hidden Markov models - Decoding problem



- Computational time of the decoding algorithm - $O\left(c^{2} T\right)$
- However, computed path might be invalid


## Hidden Markov Models

## First order hidden Markov models - Learning problem

- Determine the model parameters $p_{i j}$ and $b_{j k}$
- Given: Training sample of observed values $V^{T}$
- No method known to obtain the optimal or most likely set of parameters from the data
- However, we can nearly always determine a good solution by the forward-backward algorithm
- General expectation maximisation algorithm
- Iteratively update weights in order to better explain the observed training sequences


## Hidden Markov Models

First order hidden Markov models - Learning problem

- Probability that the model is in state $\omega_{i}(t)$ and will generate the remainder of the given target sequence:

$$
\beta_{i}(t)= \begin{cases}0 & t=T \text { and } \omega_{i}(t) \neq \omega_{0}  \tag{31}\\ 1 & t=T \text { and } \omega_{i}(t)=\omega_{0} \\ \sum_{j} \beta_{j}(t+1) p_{i j} b_{j k} v(t+1) & \text { otherwise }\end{cases}
$$

## Hidden Markov Models

## First order hidden Markov models - Learning problem

- $\alpha_{i}(t)$ and $\beta_{i}(t)$ only estimates of their true values since transition probabilities $p_{i j}, b_{j k}$ unknown
- Probability of transition between $\omega_{i}(t-1)$ and $\omega_{j}(t)$ can be estimated
- Provided that the model generated the entire training sequence $V^{T}$ by any path

$$
\begin{equation*}
\gamma_{i j}(t)=\frac{\alpha(t-1) p_{i j} b_{j k} \beta_{j}(t)}{P\left(V^{T} \mid \Theta\right)} \tag{32}
\end{equation*}
$$

- Probability that model generated sequence $V^{T}$ :

$$
\begin{equation*}
P\left(V^{\top} \mid \Theta\right) \tag{33}
\end{equation*}
$$

## Hidden Markov Models

First order hidden Markov models - Learning problem

- Calculate improved estimate for $p_{i j}$ and $b_{j k}$

$$
\begin{gather*}
\overline{p_{i j}}=\frac{\sum_{t=1}^{T} \gamma_{i j}(t)}{\sum_{t=1}^{T} \sum_{k} \gamma_{i k}(t)}  \tag{34}\\
\overline{b_{j k}}=\frac{\sum_{t=1, v(t)=v_{k}}^{T} \sum_{l} \gamma_{j l}(t)}{\sum_{t=1}^{T} \sum_{l} \gamma_{j l}(t)} \tag{35}
\end{gather*}
$$

- Start with rough estimates of $p_{i j}$ and $b_{j k}$
- Calculate improved estimates
- Repeat until some conversion is reached


## Hidden Markov Models

First order hidden Markov models - Lerning problem

## Forward-Backward algorithm

1 initialise $p_{i j}, b_{j k}, V^{\top}$, convergence criterion $\Theta, z \leftarrow 0$
$2 \quad$ do $z \leftarrow z+1$
$\begin{array}{ll}3 & \text { compute } \overline{p_{i j}(z)} \\ 4 & \text { compute } \overline{b_{j k}(z)} \\ 5 & p_{i j}(z) \leftarrow \overline{p_{i j}(z)}\end{array}$
6
$b_{j k}(z) \leftarrow b_{j k}(z)$
7 until $\max _{i, j, k}\left[p_{i j}(z)-p_{i j}(z-1), b_{j k}(z)-b_{j k}(z-1)\right]<\Theta$ (convergence achieved)
8 return $p_{i j} \leftarrow p_{i j}(z), b_{j k} \leftarrow b_{j k}(z)$
9 end

## Hidden Markov Models

First order hidden Markov models

- Context prediction with hidden Markov models:
(1) Given the transition model, estimate all $p_{i j}$ and $b_{j k}$
(2) Given a sequence $V^{T}$, decode the most probable sequence of hidden states
( Extrapolate the sequence of expected hidden states


## Outline

## Markov prediction approaches

(1) Introduction and Markov properties
(2) Markov chains
(3) Hidden Markov Models

4 Context prediction with Markov approaches

- Properties of Markov prediction approaches
(5) Conditional random fields
- Context prediction with CRF
- Properties of CRF prediction approaches
(6) Conclusion


## Context prediction with Markov approaches

- Given: Sequence of contexts $\xi_{0-k+1}, \ldots, \xi_{0}$
- Generate Markov chain representing the transition probabilities for each pair of observations
- Now possible: Provide probability distribution on the next outcome
- Can also be generalised to higher order Markov processes
- Several iterations of this process provide higher prediction horizons


## Properties of Markov prediction approaches

## Memory and processing load

- Runtime dependent on size of probability graph $G$
- C: Set of different context values
- The number of states of the Markov chain: $C$.
- Time to find most probable next state is $O(|C|)$ in the worst case.
- Every arc to possible following context to be considered.
- $|C|-1$ arcs existent in the worst case.
- Most probable $n$ future context time series elements
- Naive computation time: $O\left(|C|^{n}\right)$
- When transition probabilities stored to a matrix, one matrix multiplication for every future context
- Computation time: $O\left(n \cdot|C|^{2}\right)$.


## Properties of Markov prediction approaches

## Memory and processing load

- Memory requirements
- Dependent on the number of contexts observed - size of the transition matrix
- Order 1: $O\left(|C|^{2}\right)$
- Order k: $O\left(|C|^{k+1}\right)$


## Properties of Markov prediction approaches

## Prediction horizon

- Prediction horizon can be extended by iterative prediction
- Utilise predicted contexts as input
- Problem: Less accurate
- Predicted contexts more error prone than measured values


## Properties of Markov prediction approaches

## Adaptability

- The Markov prediction approach is well able to adapt to changing environments
- Adapt context transition probabilities
- Consideration of new events possible
- Rebuild of transition matrix required


## Properties of Markov prediction approaches

## Multi-dimensional time series

- The Markov prediction algorithm is not suited for multi-dimensional time series
- Designed for one-dimensional Input
- Possible: Aggregation of multi-dimensional time series to one-dimensional time series.


## Properties of Markov prediction approaches

## Iterative prediction

- Iterative Prediction possible
- Steep decrease in prediction accuracy expected since prediction horizon is only 1
- Increase of prediction horizon possible by Aggregation of context sequence of fixed length in one Markov state
- Prediction horizon fixed
- Increase in Memory consumption and processing time
- When I contexts are aggregated: $I^{C}$ states
- Runtime:

$$
O\left(n \cdot I^{C^{2}}\right)
$$

- Memory consumption:

$$
\begin{aligned}
& O\left(I^{C^{2}}\right) \text { (order one) } \\
& O\left(I^{C^{k+1}}\right) \text { (order k) }
\end{aligned}
$$

## Properties of Markov prediction approaches

## Prediction of context durations

- Prediction of context duration not possible
- Only simple sequence of occurring contexts possible


## Properties of Markov prediction approaches

## Approximate matching of patterns

- Exact pattern matching
- The Markov prediction algorithm utilises exact pattern matching


## Properties of Markov prediction approaches

## Context data types

- All context data types supported
- Every distinct context type one state
- Probably drastic increase in runtime and memory consumption for numeric context types
- Possible: Assign intervals to states


## Properties of Markov prediction approaches

## Pre-processing

- Pre-processing required to construct context transition probabilities
- On-line approach feasible - learning
- Runtime: $O(k)$
- Count frequency of specific context transitions in training time series of length $k$


## Aspects of prediction algorithms

## Summary

|  | IPAM | ONISI | Markov | CRF |
| :--- | :---: | :---: | :---: | :---: |
| Numeric Contexts | yes | no | yes |  |
| Non-numeric Contexts | yes | yes | yes |  |
| Complexity | $O(k)$ | () | $O\left(C^{2}\right)$ |  |
| Learning ability | (no) | yes | yes |  |
| Approximate matching | no | no | no |  |
| Multi-dim. TS | (no) | (no) | (no) |  |
| Discrete data | yes | yes | yes |  |
| Variable length patterns | no | yes | no |  |
| Multi-type TS | yes | no | (no) |  |
| Continuous data | no | no | no |  |
| Pre-processing | $O(k)$ | - | $O(k)$ |  |
| Context durations | no | no | no |  |
| Continuous time | no | no | yes |  |

## Properties of Markov prediction approaches

## Conclusion

- Markov processes are straightforward and easily applied to context prediction tasks.
- Model can be applied to numerical and non-numerical data alike.
- Prediction that reaches farther into future implicitly utilises already predicted data which might consequently decrease the prediction accuracy.


## Outline

## Markov prediction approaches

(1) Introduction and Markov properties
(2) Markov chains
(3) Hidden Markov Models

4 Context prediction with Markov approaches

- Properties of Markov prediction approaches
(5) Conditional random fields
- Context prediction with CRF
- Properties of CRF prediction approaches
(6) Conclusion


## Conditional random fields <br> Introduction

- Undirected graphical model ${ }^{67}$
- Similar to HMM
- HMM specific CRF
- Relax assumptions about input and output sequence
- Instead of constant transition probability: Arbitrary functions that vary across positions in sequence of hidden states
- Vertices represent random variables
- Edges represent dependency between two random variables

[^4]
## Conditional random fields

## Introduction

- Layout of inner-graph (hidden states) arbitrary
- Input sequence: $X$
- Inner states: $Y$
- Conditional dependency of each $Y_{i}$ on $X$ defined through set of feature functions

$$
\begin{equation*}
f\left(i, Y_{i-1}, Y_{i}, X\right) \tag{36}
\end{equation*}
$$

- Each feature assigned numerical weight


## Conditional random fields

## Training

- Various learning approaches to train CRF
- Gradient based
- Quasi-Newton-approach
- Sequences provided of which also desired output is known
- CRF parameters are adapted to match a maximum number of training sequences


## Conditional random fields

## Applications

- Applications similar to HMM
- Labeling or parsing of sequential data
- Natural language text
- Biological sequences
- Classification of proteins
- Prediction of the secondary structure of DNS and proteine
- Image recognition and image resauration


## Conditional random fields

## Discussion

- Generative models
- HMMs, stochastic grammars, ...
- Assign joint probability to paired observations
- Discriminative models
- Maximum entropy Markov models, Conditional random fields,


## Conditional random fields

## Discussion

- Problem of generative models
- To define joint probability over observation sequences (e.g. words or nucleotides), all possible observation sequences are enumerated
- Not practical
- Multiple interacting features
- Long range dependencies
- Conditional probability depends on fixed, dependent features
- Instead of arbitrary independent features
- Very strict independence assumptions on observations
- e.g. conditional independence


## Conditional random fields

## The label bias problem



- Problem of (classical) discriminative models (e.g. MEMM)
- Label bias problem ${ }^{8}$
- Conservation of score mass
- States with fewer outgoing transitions are tendentially biased
- States with low-entropy next state distributions will take little notice of observations

[^5]
## Conditional random fields

## The label bias problem



- Solutions proposed to solve the label bias problem
- Change state transition structure of the model
- Collapse states (here: 1 and 4)
- Delay branching until discriminating observation
- Start with fully connected model
- Let training figure out good structure


## Conditional random fields

## The label bias problem



- Problems with the solutions proposed
- Change state transition structure of the model
- Not always possible
- May lead to combinatorical explosion
- Start with fully connected model
- Preludes use of valuable prior structural knowledge


## Conditional random fields

## The label bias problem



- Requirement for proper solutions
- Model that accounts for whole state sequences at once
- Let transitions 'vote' more strongly than others
- Score mass will not be conserved
- Transitions can amplify or dampen received mass


## Conditional random fields

## Discussion

- Conditional random fields
- Solve label bias problem
- Single exponential model for joint probability of sequences
- Less impacted by higher-order dependencies between states


## Conditional random fields

## Algorithmic model

## Definition: Conditional random field

Let $G=(V, E)$ be a graph such that $Y=\left(Y_{v}\right)_{v \in V}$. Then $(X, Y)$ is a conditional random field when the random variables $Y_{v}$ obey the Markov property with respect to the graph:
$p\left(Y_{v} \mid X, Y_{w}, w \neq v\right)=p\left(Y_{v} \mid X, Y_{w}, w \sim v\right)$

- $w \sim v$ means that $w$ and $v$ are neighbours in $G$
- A CRF is a random field conditioned on the observation $X$
- In general, the graphical structures of $X$ and $Y$ are not the same.


## Conditional random fields

## Random field

- Random field
- Generalization of a stochastic process
- Underlying parameter need no longer be a simple real
- Can instead be multidimensional vector space


## Conditional random fields

## Random field

## Random field

Let $S=X_{1}, \ldots, X_{n}$, with the $X_{i}$ in $\{0,1, \ldots, G-1\}$ being a set of random variables on the sample space $\Omega=\{0,1, \ldots, G-1\}^{n}$. A probability measure $\pi$ is a random field if, for all $\omega$ in $\Omega, \pi(\omega)>0$.

## Conditional random fields

## Example: HMM-like CRF

$$
\begin{align*}
f_{y^{\prime}, y}\left(\langle u, v\rangle,\left.y\right|_{<u, v>}, x\right) & =\delta\left(y_{u}, y^{\prime}\right) \delta\left(y_{v}, y\right)  \tag{37}\\
g_{y, x}\left(v,\left.y\right|_{v}, x\right) & =\delta\left(y_{v}, y\right) \delta\left(x_{v}, x\right) \tag{38}
\end{align*}
$$

- Feature for each state pair $\left(y, y^{\prime}\right)$ and each state-observation pair $(y, x)$


## Conditional random fields

## Example: HMM-like CRF

- CRF more expressive: Can model more cases than HMM
- Features do not need to specify completely a state or observation.
- Therefore, model can be estimated from less training data
- CRFs share all convexity properties of general maximum entropy models


## Conditional random fields

## Example: HMM-like CRF



- Graphical structures of HMMs and CRFs
- Circle indicates that variable is not generated by the model


## Conditional random fields

## Example: HMM-like CRF



- CRF: Entire observation sequence combined


## Conditional random fields

## Experiments

| model | error | oov error |
| ---: | :---: | :---: |
| HMM | $5.69 \%$ | $45.99 \%$ |
| MEMM | $6.37 \%$ | $54.61 \%$ |
| CRF | $5.55 \%$ | $48.05 \%$ |
| MEMM $^{+}$ | $4.81 \%$ | $26.99 \%$ |
| CRF $^{+}$ | $4.27 \%$ | $23.76 \%$ |
| ${ }^{+}$Using spelling features |  |  |

## Conditional random fields

## Experiments



- Seeker robot (black) tries to tag on of the other players
- Simplified variant: Non-seeker robots follow hourclass pattern


## Conditional random fields

## Experiments

- CRF and HMM accuracy for identifying the seeker

|  | Hourglass |  |  | Unconstrained |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Features | HMM Acc. | CRF Acc. | $\ell(Y \mid X)$ | HMM Acc. | CRF Acc. | $\ell(Y \mid X)$ |
| Positions | 33.1 | 53.6 | -959.7 | 37.1 | 37.8 | -1354.5 |
| Velocities | 68.4 | 89.4 | -717.1 | 55.7 | 70.4 | -1206.5 |
| Velocity Thresholds |  |  |  |  |  |  |
| W $=\frac{1}{60}$ th sec. | 62.5 | 71.2 | -818.0 | 46.8 | 58.6 | -1148.6 |
| $\mathrm{~W}=0.1 \mathrm{sec}$. | 63.0 | 73.9 | -784.3 | 46.0 | 62.4 | -1099.2 |
| $\mathrm{~W}=0.5 \mathrm{sec}$ | 63.6 | 80.6 | -708.8 | 68.9 | 71.9 | -983.1 |
| $\mathrm{~W}=1.0 \mathrm{sec}$ | 60.2 | 83.1 | -721.8 | 67.8 | 75.3 | -980.9 |
| $\mathrm{~W}=1.5 \mathrm{sec}$ | 56.9 | 85.5 | -731.7 | 68.8 | 77.8 | -1004.7 |
| $\mathrm{~W}=2.0 \mathrm{sec}$. | 53.7 | 87.1 | -751.1 | 72.1 | 77.3 | -1036.3 |
| Chasing | 75.8 | 95.4 | -622.3 | 65.5 | 80.4 | -1058.3 |
| Distance (U) | 46.6 | 49.5 | -779.7 | 43.5 | 42.3 | -604.4 |
| Distance (N) | 46.6 | 49.9 | -200.6 | 43.5 | 58.0 | -223.4 |
| Distance \& Chasing (U) | 75.6 | 99.3 | -90.6 | 65.8 | 93.9 | -181.8 |
| Distance \& Chasing (N) | 75.6 | 99.3 | -115.3 | 65.8 | 97.6 | -112.2 |
| Many Features | 72.4 | 98.1 | -172.2 | 63.4 | 98.5 | -178.9 |
| Redundant Features | 72.4 | 95.7 | -509.3 | 52.7 | 93.8 | -6432.3 |

## Context prediction with CRF

## Prediction procedure

- Context prediction with CRF:
(1) Given the transition model, estimate all probabilites between states and state action probabilites
(2) Given a sequence $V^{T}$, decode the most probable sequence of hidden states
( Extrapolate the sequence of expected hidden states


## Aspects of prediction algorithms

## Summary

|  | IPAM | ONISI | Markov | CRF |
| :--- | :---: | :---: | :---: | :---: |
| Numeric Contexts | yes | no | no | no |
| Non-numeric Contexts | yes | yes | yes | yes |
| Complexity | $O(k)$ | ()$^{2}$ | $O\left(C^{2}\right)$ | $O\left(C^{2}\right)$ |
| Learning ability | (no) | yes | yes | yes |
| Approximate matching | no | no | no | no |
| Multi-dim. TS | (no) | (no) | (no) | (no) |
| Discrete data | yes | yes | yes | yes |
| Variable length patterns | no | yes | no | (yes) |
| Multi-type TS | yes | no | (no) | (no) |
| Continuous data | no | no | no | no |
| Pre-processing | $O(k)$ | - | $O(k)$ | $O(k)$ |
| Context durations | no | no | no | no |
| Continuous time | no | no | yes | yes |

## Properties of CRF prediction approaches

## Conclusion

- CRF processes are straightforward and easily applied to context prediction tasks.
- Model can be applied to numerical and non-numerical data alike.
- Prediction that reaches farther into future implicitly utilises already predicted data which might consequently decrease the prediction accuracy.


## Conclusion


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